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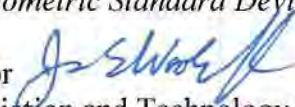
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OFFICE OF
SOLID WASTE AND EMERGENCY
RESPONSE

OSWER 9200.2-82

MEMORANDUM

SUBJECT: Transmittal of *Update of the Adult Lead Methodology's Default Baseline Blood Lead Concentration and Geometric Standard Deviation Parameters*

FROM: James E. Woolford, Director 
Office of Superfund Remediation and Technology Innovation

TO: Superfund National Policy Managers, Regions 1–10
Regional Risk Leads, Regions 1–10

Purpose

The purpose of this memorandum is to transmit the document, *Update of the Adult Lead Methodology's Default Baseline Blood Lead Concentration and Geometric Standard Deviation Parameters*. This guidance document provides the technical basis for updating the default baseline blood lead concentration and default geometric standard deviation input parameters of the Adult Lead Methodology. This document is primarily intended for Regional risk assessors.

Background

The Adult Lead Methodology (ALM) is used to assess lead risks from the soil at non-residential Superfund sites. The baseline blood lead concentration input parameter of the ALM represents the geometric mean blood lead concentration in women of child-bearing age and the geometric standard deviation (GSD) input parameter is a measure of the inter-individual variability in these concentrations.

Default values for these input parameters were originally derived from an analysis of blood lead data for U.S. women 17–45 years of age, from Phase 1 (1988 to 1991) of the Third National Health and Nutrition Examination Survey (NHANES III) as well as consideration of available site-specific data on blood lead concentrations and GSDs. EPA prepared updated estimates for these two parameters in 2002, using data from Phase 1 and 2 (1988 to 1994) of

estimates for these two parameters in 2002, using data from Phase 1 and 2 (1988 to 1994) of NHANES III. The purpose of this report is to provide updated estimates for these parameters using data from the NHANES surveys that were conducted from 1999–2004.

Implementation

This document provides updated values for the default blood lead concentration and the geometric standard deviation input parameters of the Adult Lead Methodology. However, recent scientific evidence has demonstrated adverse health effects at blood lead concentrations below 10 µg/dL down to 5 µg/dL, and possibly below. OSRTI is developing a new soil lead policy to address this new information. Until that soil lead policy is finalized, regional risk assessors and managers should consult with the Lead Committee of the Technical Review Workgroup for Metals and Asbestos (TRW) before applying these updated values for risk assessment.

If you have any questions, please contact me or have your staff contact Aaron Yeow at yeow.aaron@epa.gov or Michael Beringer at beringer.michael@epa.gov.

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**OSWER 9200.2-82
JUNE 2009**



**UPDATE OF THE ADULT LEAD METHODOLOGY'S
DEFAULT BASELINE BLOOD LEAD CONCENTRATION
AND GEOMETRIC STANDARD DEVIATION PARAMETERS**

**PREPARED BY THE
LEAD COMMITTEE OF THE
TECHNICAL REVIEW WORKGROUP FOR METALS AND ASBESTOS**

**OFFICE OF SUPERFUND REMEDIATION AND TECHNOLOGY INNOVATION
UNITED STATES ENVIRONMENTAL PROTECTION AGENCY**

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1.0 Introduction

In 1996 the Technical Review Workgroup for Lead (TRW) recommended the use of the Adult Lead Methodology (ALM) (U.S. EPA, 1996) for assessing risks to adults from exposures to lead in soil at non-residential Superfund sites. The background blood lead concentration ($PbB_{adult,0}$) parameter in the ALM represents the geometric mean blood lead concentration (PbB) ($\mu\text{g/dL}$) in women of child-bearing age, in the absence of exposures at the site being assessed. The geometric standard deviation parameter ($GSD_{i,adult}$) is a measure of the inter-individual variability in blood lead concentrations in a population whose members are exposed to the same non-residential environmental lead levels. Default values for both $PbB_{adult,0}$ and $GSD_{i,adult}$ were originally derived from an analysis of blood lead data for U.S. women 17–45 years of age, from Phase 1 (1988 to 1991) of the Third National Health and Nutrition Examination Survey (NHANES III)¹ as well as consideration of available site-specific data on PbBs and GSDs (U.S. EPA, 1996). The TRW prepared updated estimates for these two parameters in 2002, using data from Phase 1 and 2 (1988 to 1994) of NHANES III (U.S. EPA, 2002). The purpose of this report is to provide updated estimates for $PbB_{adult,0}$ and $GSD_{i,adult}$ using data from the NHANES surveys that were conducted from 1999–2004. Although the Centers for Disease Control (CDC) releases data from the continuous NHANES in 2-year increments, it is recommended to use four or more years of data when estimating parameters for demographic sub-domains (CDC, 2005a).

This document provides the technical basis for updating the $PbB_{adult,0}$ (GM) and $GSD_{i,adult}$ (GSD) parameters and details on how the updated estimates for the parameters were calculated. The intended audience for this document is risk assessors who are familiar with using the ALM. For background and further detail on the use of the ALM in Superfund lead risk assessment, please refer to U.S. EPA (1996) and the TRW Lead Committee website (<http://www.epa.gov/superfund/lead>).

2.0 Technical Approach

Information on PbB for non-institutionalized U.S. women 17–45 years of age was extracted from the NHANES database (CDC, 2005b). Data from three 2-year cycles of the continuous NHANES (1999–2004) were used in this analysis in accordance with CDC recommendations (CDC, 2005a). Results reported at less than the detection limit of $0.3 \mu\text{g/dL}$ were analyzed by a variety of methods, including $\frac{1}{2}$ the detection limit (which is consistent with the 2002 analysis [U.S. EPA, 2002] and EPA guidance [U.S. EPA, 1998]), and other statistical methods such as the Cohen maximum likelihood estimation method (Cohen, 1959), the Tobit maximum likelihood estimation method (Tobin, 1958), regression on order statistics (Newman et al., 1989) and the Kaplan-Meier method (Kaplan and Meier, 1958)². Further discussion of these methods is presented in the Appendix. Estimates of the $PbB_{adult,0}$ and $GSD_{i,adult}$ were calculated using SAS[®] software, Version 9.1 of the SAS System for Microsoft Windows³ and the sample weights recommended by CDC (2005a).

¹ For more information see http://www.cdc.gov/nchs/about/major/nhanes/nhanes2005-2006/faqs05_06.htm#question%2010 or EPA's National Center for Environmental Assessment's [Handbook for Use of Data from the National Health and Nutrition Examination Surveys \(NHANES\): A Goldmine of Data for Environmental Health Analyses](#)

² This method is equivalent to the Kaplan-Meier approach in EPA ProUCL Software.

³ SAS and all other SAS Institute Inc. product or service names are registered trademarks or trademarks of SAS Institute Inc. in the USA and other countries. Although a newer version is available, it is unlikely the results would be affected.

Standard errors for the estimates of the GM were estimated using the Taylor linearization method in SUDAAN⁴ (Version 9.0, SAS-callable version). SUDAAN is designed to compute statistics (*e.g.*, means and percentiles) and their standard errors for data that are from complex sample surveys such as the NHANES. SUDAAN does not calculate estimates of population variance, such as the GSD. To the best of our knowledge, a Taylor linearization approach for estimating the standard error is not available for the GSD; therefore, standard errors for the GSD were estimated using a SAS macro³ that implements a jackknife method. Parameter estimates used the sample weights provided in the NHANES demographic data files (CDC, 2005a). Standard errors for the GSD were estimated using the sample weights and the masked variance units (*i.e.*, pseudo-strata and pseudo-primary sampling units which are also provided in the NHANES demographic files). The sample weights account for the unequal probabilities of selection of survey participants, the non-response of some participants, and are adjusted to population controls. The masked-variance units account for the multistage sampling design and are necessary to estimate accurate standard errors for parameter estimates.

3.0 Results

The 1999–2004 NHANES data provided 4,589 blood lead measurements for non-institutionalized U.S. women 17–45 years of age. GM PbB and GSD were derived from the data using a variety of methods. Table 1 presents the estimates for the GM PbB and GSD calculated by the various methods.

Table 1. Results of various methods for determining geometric mean baseline blood lead concentration (GM) and geometric standard deviation (GSD) for U.S. Women age 17–45 years for NHANES III (1999-2004)		
Method	GM (µg/dL)	GSD
Assigning ½ detection limit (0.15) to non-detects	1.0	1.8
Cohen MLE ^a	1.0	1.8
Tobit MLE ^b	1.0	1.8
ROS ^c	1.0	1.7
Helsel Kaplan-Meier ^d	1.0	1.8

^aCohen: Maximum likelihood estimation method of Cohen (1959).

^bTobit: Maximum likelihood estimation method by Tobit regression (Tobin, 1958).

^cROS: Regression on Order Statistics method (Newman et al., 1989).

^dHelsel Kaplan-Meier: Kaplan-Meier estimate calculated with Helsel’s KMStats.xls spreadsheet with Efron’s bias correction (Kaplan and Meier, 1958).

⁴ SUDAAN® is a registered trademark of the Research Triangle Institute. © 2005 Research Triangle Institute. All rights reserved. Although a newer version is available, it is unlikely the results affected.

³ The SAS macro implements the ‘leave one out’ jackknife method (*e.g.*, Research Triangle Institute. 2004. SUDAAN Language Manual, Release 9.0. Research Triangle Park, NC).

The GM PbB is estimated to be 1.0 µg/dL, and the GSD is estimated to be 1.8. Table 2 presents the updated estimates as well as the estimates from the previous analyses.

Table 2. Geometric mean baseline blood lead concentration (GM, µg/dL) and GSD estimates of U.S. women, 17–45 years of age for NHANES III (1999–2004)			
NHANES Data	<i>n</i>	GM (µg/dL)	GSD
1988–1991 NHANES III – EPA ALM (1996)	–	1.7–2.2	1.8–2.1
1988–1994 NHANES III – EPA 2002 Update	5,016	1.5	2.1
1999–2004 NHANES – Current Update	4,589	1.0	1.8

4.0 Uncertainty

There are two main sources of uncertainty in this analysis: the unknown PbB concentrations for data reported as non-detects and the potential bias in the estimates, particularly the GSD, that can occur when a few observations have an undue influence on the estimate due to large sample weights.

As blood lead levels continue to decline in the U.S. population, the number of non-detects in the NHANES data has the potential to become an important source of uncertainty in estimates of PbB and GSD. However, the detection limit for measuring lead concentration in blood was lowered from 1.0 µg/dL for the 1988–1994 NHANES III to 0.3 µg/dL for the 1999–2004 NHANES. The lower detection limit removes a considerable source of uncertainty that was present in previous estimates of the GM (U.S. EPA, 2002) as the rate of non-detects in the NHANES 1999–2004 data (~2%) is much lower than the rate of non-detects in the 1988–1994 NHANES III data (~21%). Nonetheless, the potential effect of the non-detects on the robustness of the estimates was explored using a variety of methods, ranging from simple substitution to more complex statistical methods (see Appendix for additional information).

An unbiased estimate for the GM can be made using any subset of the PbB concentrations by using the sample weights included in the NHANES database. However, weighted estimates of population variability, such as the GSD, have the potential to be unduly influenced by observations that receive large sample weights. This source of uncertainty in the estimate of the GSD is partially addressed by estimating the GSD using a regression approach that is described in the *Uncertainty* section.

The results of the uncertainty analysis yielded similar estimates of the GM and GSD regardless of the method used to treat the non-detects (see Table 1). This increases confidence in the estimates and indicates that the non-detects and sample weights do not have a substantial effect on the estimates of the GM and GSD.

5.0 Recommendations

Consistent with the 2002 report (U.S. EPA, 2002), estimates of the $PbB_{adult,0}$ and $GSD_{i,adult}$ are provided for the population of non-institutionalized U.S. women 17–45 years of age. Unlike the earlier analysis, the TRW recommends using a single national estimate instead of the regional or ethnic alternatives. Feedback from Regional risk assessors indicates that the regional and ethnic information are not useful because populations move between regions and exposure is not typically ethnically homogenous. Based on this analysis of the NHANES 1999–2004 data, the updated values for the $PbB_{adult,0}$ and $GSD_{i,adult}$ parameters, 1.0 $\mu\text{g/dL}$ and 1.8, respectively, are recommended for all applications of the ALM where current and future use scenarios are assessed (see Table 3). These estimates have been shown to be robust to the two sources of uncertainties addressed in the Uncertainty Section and the Appendix.

Table 3. Recommended baseline blood lead concentration ($\mu\text{g/dL}$) and GSD estimates of U.S. women, 17–45 years of age, between NHANES III and 1999–2004 NHANES data			
	<i>n</i>	GM ($\mu\text{g/dL}$)	GSD
1999–2004 NHANES	4,589	1.0	1.8

6.0 References

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Appendix: Uncertainty Analysis

This Appendix provides additional details on the methods that were used to assess the two sources of uncertainty that could have an effect on the reliability of the estimates of the GM and GSD:

- The unknown PbB concentrations for the data reported as non-detects; and
- The potential bias in the estimates, particularly the GSD, which can occur when a few observations have an undue influence on the estimate due to large sample weights.

Several methods were used to assess the robustness of the estimates to the two sources of uncertainty described above, ranging from simple to more complex statistical methods. A search of the literature on more advanced statistical methods (*i.e.*, rather than simple substitution) for estimating the mean and standard deviation with data that include non-detects found maximum likelihood estimation (MLE) and regression on order statistics (ROS) methods to be the most often recommended. In addition to these parametric methods, Helsel (2005) recommends the use of the non-parametric Kaplan-Meier method for data that are multiply censored (*e.g.*, more than one detection limit) and a robust version of the ROS method. The use of more than one method builds confidence in the estimates when the methods produce similar estimates.

The Kaplan-Meier method has a small positive bias when the smallest value is censored, which is the case with the NHANES 1999-2004 PbB data (Helsel, 2005). Efron's bias correction was applied to address this bias. The Kaplan-Meier method is not well suited for this analysis because it is designed for data with multiple censors. The 1999-2004 NHANES data only has one detection limit, and applying the Kaplan-Meier method to singly-censored is equivalent to replacing the non-detects with the detection limit (Helsel, 2005). Nonetheless, this method was included in the analysis at the suggestion of several Regional Risk Assessors.

Helsel's robust regression on order statistics (ROS) method was not used in this analysis because the advantage of this method over the common form of the ROS method is to reduce the bias that is introduced when estimates of the arithmetic mean and standard deviation in log-space are required in the original measurement scale. Estimation of the geometric mean and geometric standard deviation are not susceptible to this bias because, by definition, the geometric mean and geometric standard deviation equal the anti-log of the mean and standard deviation of the log-transformed data.

Based on the literature search, the MLE method of Cohen (1959) was selected because of the large sample size available for this analysis ($n = 4,589$) and the good fit of the normal distribution to the log-transformed data (Figure A-1). Furthermore, MLE methods are more precise (provide narrow confidence intervals relative to other methods) and the Cohen method could be implemented with the complex survey data of NHANES using commercially available software (SAS). While MLE methods are not unbiased, the bias is a practical concern only with small sample sizes. Cohen recommends his method when samples sizes are 10 or "... slightly larger..." (Cohen, 1959); Newman et al. (1995) suggest the bias is important when the sample size is 20 or less and Helsel (2005) recommends MLE methods for sample sizes of 50 or more. Simulation tests have shown the Cohen MLE method produces estimates with lower bias and higher precision than other methods for large sample sizes (Haas and Scheff, 1990; Newman et al., 1989; Shumway, et al., 2002).

The formulas that are used to calculate the MLEs for the mean and standard deviation of the log-transformed data are shown below (Equations 1 and 2, respectively).

$$\hat{\mu} = \bar{x} - \hat{\sigma} \left(\frac{k}{n-k} \right) \left(\frac{f(\varepsilon)}{F(\varepsilon)} \right) \quad \text{Equation 1}$$

$$\hat{\sigma}^2 = \frac{s^2 + (\bar{x} - \hat{\mu})^2}{1 + \varepsilon \left(\frac{k}{n-k} \right) \left(\frac{f(\varepsilon)}{F(\varepsilon)} \right)} \quad \text{Equation 2}$$

Where:

$$\varepsilon = (DL - \hat{\mu}) / \hat{\sigma} = -1.160$$

DL = detection limit (in log units)

k = number of non-detects

n = sample size

$\hat{\mu}$ = MLE estimate of the population mean

$\hat{\sigma}^2$ = MLE estimate of the population variance

\bar{x} = sample average

s^2 = sample variance

$f(\varepsilon)$ = normal density function evaluated at ε

$F(\varepsilon)$ = cumulative normal distribution function evaluated at ε

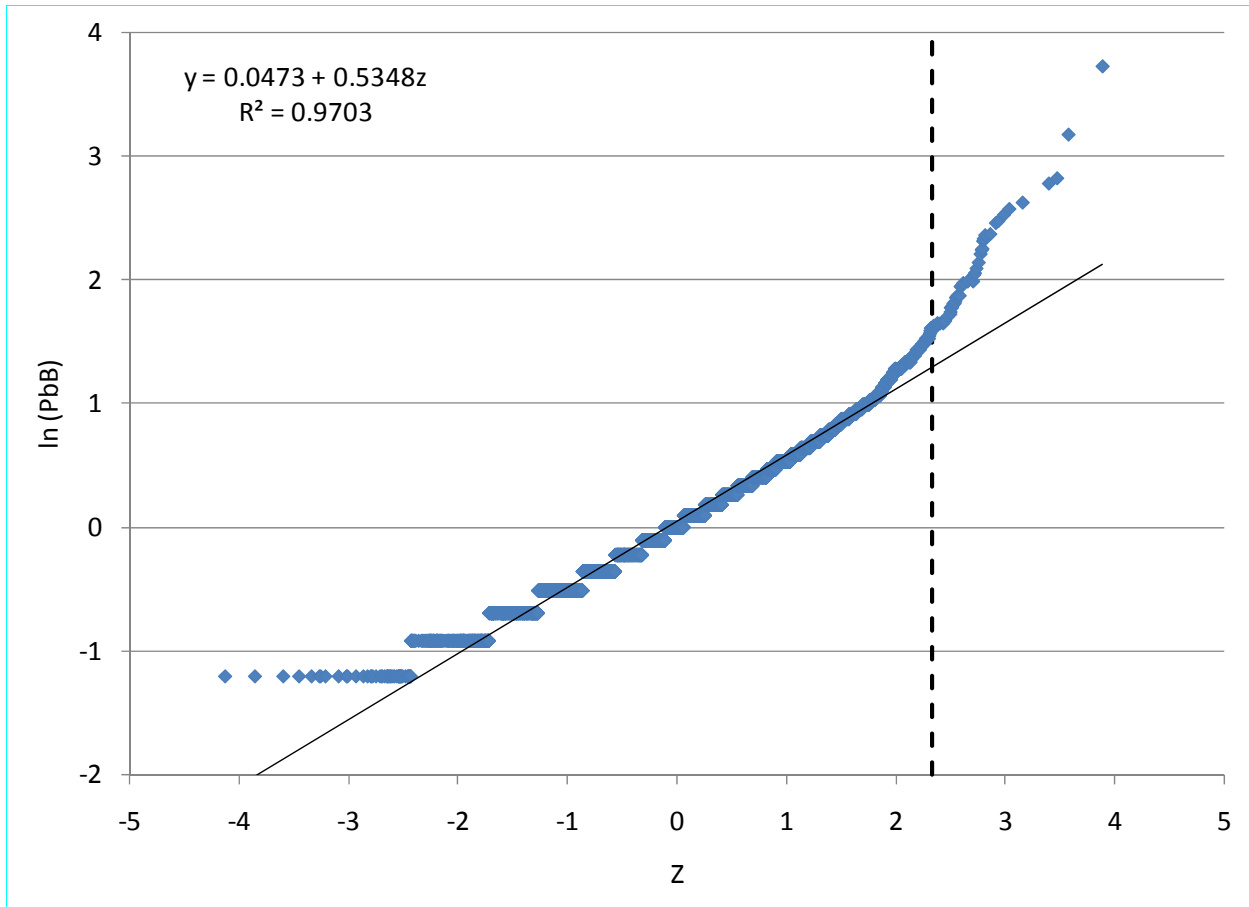


Figure A-1. Probability plot for blood lead concentration for U.S. women, 17–45 years old using only observations that are equal to or above the level of detection (0.3 µg/dL). The vertical axis is the natural log of blood lead concentration in µg/dL, the horizontal axis is the normal score (z). The vertical dotted line shows where the 99th percentile intersects with the straight line. A lognormal distribution provides a reasonably good fit to the estimated distribution, up to approximately the 99th percentile. This plot also serves as a graphical description of the Newman et al. (1989) regression on order statistics method.

These equations are typically solved iteratively (*e.g.*, Newman et al., 1995; Shumway et al., 2002), by assuming ε can be estimated by $F^{-1}(k/n)$ (assuming ε can be estimated by the inverse of the cumulative normal distribution function evaluated at k/n , or the fraction of non-detects) (Newman et al., 1995), or by using another form of Equations 1 and 2 with tabled values for an auxiliary estimation function (θ) provided by Cohen (1959) (Equations 3 and 4).

$$\hat{\mu} = \bar{x} - \theta(\bar{x} - DL) \quad \text{Equation 3}$$

$$\hat{\sigma}^2 = s^2 + \theta(\bar{x} - DL)^2 \quad \text{Equation 4}$$

The auxiliary function (θ) considers the frequency of non-detects and the difference between the detection limit (DL) and sample average (in log units).

The MLE estimates of the GM and GSD, computed by taking the anti-log of the estimate of the population mean (Equation 3) and the anti-log of the square root of the estimate of the population variance (Equation 4), are 1.0 $\mu\text{g/dL}$ and 1.8, respectively. These results are the same (to 0.1) as the results obtained by the other methods, which increases the confidence in the estimates of the GM and GSD previously obtained (Table 1). These results show that the estimates obtained with any of the methods are robust to the uncertainty around the actual values of the non-detects.

The MLE method described above is approximate. The method assumes the data are independent (*e.g.*, equally weighted), which is not the case with the NHANES PbB data. The estimates computed using the MLE method as described above did not fully account for the NHANES sample weights. The sample weights were used to compute the sample mean and variance (\bar{x} and s^2 , respectively); however, the theory behind the adjustments to the sample mean and variance that remove the bias in the estimates is based on the assumption that the blood lead data are independent normal variables (*i.e.*, all sample weights = 1/sample size). Given the low percentage of non-detects in the PbB data (~2%), the adjustments have little effect on the estimates of the GM and GSD. A version of the MLE could be developed that fully accounts for the NHANES weights; however, based on the above results, we expect that for this data set estimates of the GM and GSD would be the same (to two significant digits), so we did not pursue this additional level of complexity.

Helsel (2005) recommends an iterative solution to the MLE over the use of the auxiliary function. An iterative solution to MLEs was computed using Tobit regression (Tobin, 1958). An advantage of the Tobit regression method is it can be used with data that include non-detects with more than one limit of detection (LOD). Similar to the Kaplan-Meier method, this advantage is not useful for this dataset. Nonetheless, it was included at the suggestion of several Regional Risk Assessors. The model was estimated using SAS Proc LifeReg, which employs a Newton-Raphson algorithm to calculate the maximum of the likelihood function. The GM and GSD were estimated as the mean and slope, respectively, of the estimated regression model. The GM and GSD estimated by the Tobit regression method (1.0 $\mu\text{g/dL}$ and 1.8, respectively) are the same (to 0.1) as those estimated by the Cohen method and similar to the other methods (Table 1).

To address the uncertainty in the parameter estimates due to the sample weights, the regression on order statistics (ROS) method recommended by Newman et al. (1989) was selected because it does not entail squaring the sample weight to compute the estimate of the GSD. Cohen (1959) and Helsel (2005) also recommend the ROS method for small sample sizes, although Helsel's recommendation is based on the assumption that estimates of the mean and standard deviation must be back-transformed to the original measurement scale, which is not an issue in this analysis.

The ROS method assumes the data are independent and from a normal distribution, but this method has been shown to be robust to departures from normality (Gilliom and Helsel, 1986; Helsel, 2005). This method was modified to fully account for the NHANES sampling weights. Under the assumption of independent observations, the ROS method regresses the log-transformed data on the normal scores. The normal scores are calculated as shown in Equation 5.⁶ This equation was modified to include the sample weights, resulting in Equation 6. Equation 5 is a special case of Equation 6 where the weights are all equal (and therefore cancel out).

⁶ The formula for calculating normal scores is from Goovaerts (1997). The differences between the normal scores obtained using Equation 3 and more frequently used formulas (*e.g.*, Blum [1958]), are trivial for large sample sizes.

$$Z(i) = F^{-1} \left[\frac{r_i}{n} - \frac{0.5}{n} \right] \quad \text{Equation 5}$$

$$Z(i) = F^{-1} \left[\sum_{i=1}^r w_i - 0.5w_i \right] \quad \text{Equation 6}$$

Where:

F^{-1} = inverse of the cumulative normal distribution function

r_i = rank of the i^{th} observation

w_i = sample weight for i^{th} observation (adjusted so they sum to 1)

n = number of observations in the sample

$Z(i)$ = normal score of the i^{th} observation

In the ROS method, the log-transformed data above the detection limit are regressed on their sample-weighted normal scores (Figure A-1). The mean and standard deviation are estimated by the intercept and slope of the regression equation, respectively. The GM and GSD are computed by taking the anti-log of the estimates of the mean (intercept) and standard deviation (slope). The estimate of the GM by the ROS method is 1.0 $\mu\text{g/dL}$; the same estimate derived by the other methods (Table 1). The GSD estimated by the ROS method, 1.7, is approximately 0.1 less than the estimates obtained by the other methods (see Table 1), indicating there is no substantial bias in the estimate caused by large sample weights.

The consistency in the estimates strongly indicates the estimates recommended in Table 3 are robust to the uncertainty of the actual values of the non-detects and indicates the potential for bias caused by a few large sample weights is not a concern.